Modeling Financial Crises Mutations

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Introduction

- → Importance of financial crises (Latin America, in Asia, Scandinavia, ERM, Russia, Asian, Lehman brothers, Greece, Ireland, Portugal...
- \rightarrow Modeling crisis is a crucial issue for the analysis of the crisis, Optimal policy set up, Early Warning Systems,...

Introduction - Modeling Methods

How to model a financial crisis?

→ **Static models** - rich literature (Kaminski et al.,1998; Berg and Patillo,1999; Kumar et al.,2003 Bussiere and Fratzscher, 2006; etc.)

\rightarrow Dynamic models

- Duration model; (Tudela, 2005)
- Markov Switching model; (Abiad et al. 2003).

Recently, Dynamic probit model:

"Currency crises early warning systems: why they should be dynamic" (2010) B.Candelon, E. Dumitrescu and C. Hurlin.

Introduction - Economics 1

What is a financial crisis?

Frederic Mishkin: Nonlinear disruption,..., so that financial markets are unable to channel funds to those with the most productive investment opportunities.

- Currency crisis.
- Banking crisis.
- Sovereign debt crisis.

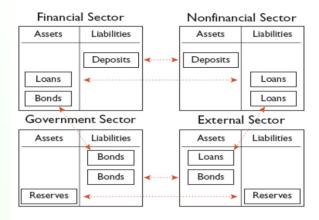
Introduction - Economics 2

From one crisis to another one, examples..

- Banking to sovereign debt in Europe. (Candelon and Palm, 2010)
- ▶ Banking to currency (twin crisis), Ecuador 1999,... (Glick and Hutchinson, 1999, Reinhart and Rogoff, 2010)

Introduction - Economics 3

... and Economic theory (Balance sheet approach)



Originality

Dynamic EWS models

- \rightarrow Dynamic model with binary crisis variables.
- ightarrow Multivariate to take into account for the potential crisis mutation

Sketch of the presentation

1. Methodology:

- The Model.
- Exact Maximum Likelihood Estimation.

2. Empirical Application:

- Data.
- Defining the crisis periods.
- Bivariate vs multivariate model.
- The Ecuador example: conditional probability and IRF.

Take Aways

- 1. MDP is a much more parcimonious model.
- 2. MDP shows better in-sample properties as it takes into account crisis mutation.
- \rightarrow It should be implemented as often as possible.

Methodology- The Model 1

$$y_{m,t}^* = \pi_{m,t} + \epsilon_{m,t}, \tag{1}$$

and

$$y_{m,t} = \mathbb{1}(y_{m,t}^* > 0),$$
 (2)

with $m \in \{c, b, s\}$, $\pi_{m,t}$ being the expected value of $y_{m,t}$ that may depend on covariates which vary across markets, country and time and with $E(\epsilon_{m,t}|\pi_{m,t})=0$, $Var(\epsilon_{m,t}|\pi_{m,t})=\Gamma$, $Cov(\epsilon_{m,t},\epsilon_{m',t}|\pi'_{m,t},\pi_{m',t'})=\omega_{mm'}$ when i=i', t=t' and zero whenever $i\neq i'$ and $t\neq t'$.

Methodology- The Model 2

$$y_{m,t}^* = \alpha_m + x'_{m,t-1}\beta_m + \sum_{m'} y_{m',t-1}\Delta_{m,m'} + \sum_{m'} \Gamma_{m,m'}\pi_{m',t-1} + \varepsilon_{m,t}$$

$$y_{m,t} = \mathbb{1}(y_{m,t}^* > 0).$$
(3)

Interpretation for persistence and causality:

- \rightarrow Persistence: diagonal terms \triangle (non-linear) and Γ (linear).
- ightarrow Causality (Granger): off-diagonal terms of Δ and Γ .

Methodology- the Model 3

Finally, the disturbances $\varepsilon_t = [\varepsilon_{c,t} \ \varepsilon_{b,t} \ \varepsilon_{s,t}]'$ are trivariate normally distributed with a 3 × 3 symmetric matrix $\tilde{\Omega}$:

$$\tilde{\Omega} = \begin{pmatrix} \sigma_c^2 & \rho_{bc}\sigma_b\sigma_c & \rho_{sc}\sigma_c\sigma_s \\ \rho_{bc}\sigma_b\sigma_c & \sigma_b^2 & \rho_{sb}\sigma_b\sigma_s \\ \rho_{sc}\sigma_c\sigma_s & \rho_{sb}\sigma_b\sigma_s & \sigma_s^2 \end{pmatrix}, \tag{4}$$

where $\rho_{m,m-1}$ represents the correlation coefficients. It is also assumed that $\tilde{\varepsilon}_t$ is i.i.d so that the covariance matrix for all T observations is given by $V(\tilde{\varepsilon}) = I_N \otimes \tilde{\Omega}$, $\tilde{\Omega}$ being a flexible covariance matrix.

Methodology- the Model 4

1.
$$\pi_t = \alpha + x'_{t-1}\beta + y'_{t-1}\Delta + \Gamma'\pi_{t-1}. \tag{5}$$

2. $\pi_t = \alpha + \mathbf{x}'_{t-1}\beta, \tag{6}$

3.
$$\pi_{t} = \alpha + x'_{t-1}\beta + y'_{t-1}\Delta, \tag{7}$$

4. $\pi_{t} = \alpha + X'_{t-1}\beta + \Gamma'\pi_{t-1}, \tag{8}$

Methodology- Exact ML 1

The FIML estimates are obtained by maximizing the log-likelihood:

$$LogL(y|z,\theta;\Omega) = \sum_{t}^{T} Log\Phi_{3,\varepsilon}(w_t; Q_t\Omega Q_t)$$
 (9)

where Q_t is a diagonal matrix whose main diagonal elements are $q_{m,t}=2y_{m,t}-1$ and thus depends on the realization or not of the events $(q_{m,t}=1 \text{ if } y_{m,t}=1 \text{ and } q_{m,t}=-1 \text{ if } y_{m,t}=0, \forall m \in \{c,b,s\})$. Besides, the elements of the vector $w_t=[w_{1,t},...,w_{3,t}]$ are given by $w_{m,t}=q_{m,t}\pi_{m,t}$.

Methodology- Exact ML 2

Ideas for the empirical procedure:

- → Huguenin, Pelgrin and Holly (2009) show in a static probit framework that simulated methods lead to bias So Exact ML.
- $\to \Phi_{3,\varepsilon}(w_t; Q_t\Omega Q_t)$ is some a simple, double and triple integrals. Triple integrale can be decompose in a non-unique way into double integrals.
- ightarrow Integrals are numerically evaluated using Gauss-Legendre Quadrature rule over bounded intervals.

Application- Data

Table 1 – Database

| Country | Bivariate model | Trivariate model |
|--------------|-------------------------------|-------------------------------|
| Argentina | February 1988 - May 2010 | December 1997 - May 2010 |
| Brazil | September 1990 - May 2010 | December 1997 - May 2010 |
| Chile | January 1989 - May 2009 | May 1999 - May 2010 |
| Colombia | February 1986 - August 2009 | December 1997 - August 2009 |
| Ecuador | January 1994 - November 2007 | December 1997 - November 2007 |
| Egypt | February 1986 - June 2009 | July 2001 - June 2009 |
| El Salvador | January 1991 - November 2008 | April 2002 - November 2008 |
| Indonesia | January 1989 - August 2009 | May 2004 - August 2009 |
| Lebanon | January 1989 - April 2010 | April 1998 - April 2010 |
| Malaysia | January 1988 - March 2010 | December 1997 - March 2010 |
| Mexico | January 1988 - May 2010 | December 1997 - May 2010 |
| Peru | January 1990 - May 2010 | December 1997 - May 2010 |
| Philippines | January 1995 - February 2008 | December 1997 - February 2008 |
| South Africa | January 1988 - August 2009 | December 1997 - August 2009 |
| Turkey | January 1988 - May 2010 | December 1997 - May 2010 |
| Venezuela | February 1986 - November 2009 | December 1997 - November 2009 |

Note: Data availability.

1- The Currency Crises

Market pressure index (MPI) KLR(1998)

$$KLRm_{n,t} = \frac{\Delta e_{n,t}}{e_{n,t}} - \frac{\sigma_e}{\sigma_r} \frac{\Delta r_{n,t}}{r_{n,t}} + \frac{\sigma_e}{\sigma_i} \Delta i_{n,t},$$
 (10)

Currency crisis variable:

$$CC_{n,t} = \begin{cases} 1, & \text{if } \mathsf{KLRm}_{n,t} > 1.5\sigma_{\mathsf{KLRm}_{n,t}} + \mu_{\mathsf{KLRm}_{n,t}} \\ 0, & \text{otherwise.} \end{cases}$$
 (11)

2- The Banking Crises

Banking pressure index (BPI) von Hagen and Ho (2004)

$$BPI_{n,t} = \frac{\Delta \gamma_{n,t}}{\sigma_{\Delta \gamma}} + \frac{\Delta r_{n,t}}{\sigma_{\Delta r}},$$
(12)

Banking crisis variable:

$$BC_{n,t} = \begin{cases} 1, & \text{if } IMP_{n,t} > P_{BPI,90,n} \\ 0, & \text{otherwise.} \end{cases}$$
 (13)

3- The Sovereign Debt Crises

Pescatori and Sy (2007) use the CDS spread.

Sovereign debt crisis variable:

$$SC_{n,t} = \begin{cases} 1, & \text{if CDSspread}_{n,t} > \textit{Kernel Threshold}_n \\ 0, & \text{otherwise.} \end{cases}$$
 (14)

Table 2 – Percentage of crisis periods

| | Bivariat | e model | Trivariate model | | | |
|--------------|-----------------|----------------|------------------|----------------|-------------|--|
| | Currency crisis | Banking crisis | Currency crisis | Banking crisis | Debt crisis | |
| Argentina | 5.13 | 8.90 | 4.00 | 6.67 | 10.0 | |
| Brazil | 3.77 | 7.19 | 0.00 | 3.33 | 2.67 | |
| Chile | 6.07 | 10.0 | 5.79 | 5.79 | 3.31 | |
| Colombia | 4.95 | 9.90 | 9.22 | 12.8 | 0.00 | |
| Ecuador | 5.73 | 9.93 | 6.67 | 10.8 | 6.67 | |
| Egypt | 6.76 | 9.96 | 4.17 | 7.30 | 7.30 | |
| El Salvador | 3.65 | 9.85 | 0.00 | 0.00 | 2.50 | |
| Indonesia | 5.30 | 9.90 | 0.00 | 14.0 | 6.25 | |
| Lebanon | 9.62 | 9.96 | 1.38 | 8.97 | 2.76 | |
| Malaysia | 3.10 | 10.0 | 4.05 | 6.08 | 4.73 | |
| Mexico | 6.50 | 9.93 | 0.00 | 9.33 | 0.00 | |
| Panama | 0.00 | 9.89 | 0.00 | 6.38 | 0.00 | |
| Peru | 4.45 | 8.22 | 0.00 | 10.7 | 0.00 | |
| Phillipines | 4.90 | 9.80 | 5.69 | 6.50 | 3.25 | |
| South Africa | 6.71 | 9.89 | 7.09 | 7.80 | 4.26 | |
| Turkey | 4.80 | 8.56 | 4.00 | 6.67 | 0.00 | |
| Venezuela | 7.33 | 10.1 | 4.17 | 7.64 | 2.78 | |

Note: A percentage of crisis superior to 5% is represented in bold.

Application- Bivariate vs trivariate DPM 1

Table 3 - Bivariate Analysis

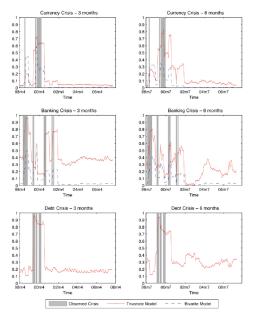
| | | | onths | 6 months | | 12 months | |
|--------------|---------------------|---|---|---|---|---|---|
| Country | | θ | Ω | θ | Ω | θ | Ω |
| Argentina | currency banking | [+ +] | $\left[\begin{array}{cc} 1 & + \\ + & 1 \end{array}\right]$ | [· +] | $\left[\begin{array}{cc} 1 & + \\ + & 1 \end{array}\right]$ | . + | |
| Chile | currency banking | $\left[\begin{array}{cc} + & \cdot \\ \cdot & \cdot \end{array}\right]$ | $\left[\begin{array}{cc} 1 & + \\ + & 1 \end{array}\right]$ | $\left[\begin{array}{cc} \cdot & \cdot \\ \cdot & \cdot \end{array}\right]$ | $\left[\begin{array}{cc} 1 & + \\ + & 1 \end{array}\right]$ | $\left[\begin{array}{cc} \cdot & - \\ \cdot & + \end{array}\right]$ | $\left[\begin{array}{cc} 1 & + \\ + & 1 \end{array}\right]$ |
| Ecuador | currency banking | $\left[\begin{array}{cc} \cdot & \cdot \\ \cdot & + \end{array}\right]$ | $\left[\begin{array}{cc} 1 & . \\ . & 1 \end{array}\right]$ | $\left[\begin{array}{cc} \cdot & \cdot \\ \cdot & + \end{array}\right]$ | $\left[\begin{array}{cc}1&.\\.&1\end{array}\right]$ | $\left[\begin{array}{cc} \cdot & \cdot \\ \cdot & + \end{array}\right]$ | $\left[\begin{array}{cc} 1 & . \\ . & 1 \end{array}\right]$ |
| Egypt | currency banking | $\left[\begin{array}{cc} + & \cdot \\ - & + \end{array}\right]$ | $\left[\begin{array}{cc} 1 & . \\ . & 1 \end{array}\right]$ | $\left[\begin{array}{cc} + & \cdot \\ - & + \end{array}\right]$ | $\left[\begin{array}{cc} 1 & - \\ - & 1 \end{array}\right]$ | $\left[\begin{array}{cc} + & \cdot \\ \cdot & + \end{array}\right]$ | $\left[\begin{array}{cc} 1 & - \\ - & 1 \end{array}\right]$ |
| Lebanon | currency banking | $\left[\begin{array}{cc} + & \cdot \\ - & + \end{array}\right]$ | $\left[\begin{array}{cc} 1 & . \\ . & 1 \end{array}\right]$ | $\left[\begin{array}{cc} \cdot & \cdot \\ \cdot & + \end{array}\right]$ | $\left[\begin{array}{cc} 1 & + \\ + & 1 \end{array}\right]$ | $\left[\begin{array}{cc} + & \cdot \\ \cdot & + \end{array}\right]$ | $\left[\begin{array}{cc} 1 & + \\ + & 1 \end{array}\right]$ |
| Mexico | currency banking | $\left[\begin{array}{cc} + & . \\ . & + \end{array}\right]$ | $\left[\begin{array}{cc}1&.\\.&1\end{array}\right]$ | + . | $\left[\begin{array}{cc}1&.\\.&1\end{array}\right]$ | $\left[\begin{array}{cc} + & . \\ . & . \end{array}\right]$ | $\left[\begin{array}{cc} 1 & + \\ + & 1 \end{array}\right]$ |
| South Africa | currency banking | $\left[\begin{array}{cc} + & \cdot \\ \cdot & + \end{array}\right]$ | $\left[\begin{array}{cc} 1 & . \\ . & 1 \end{array}\right]$ | $\left[\begin{array}{cc} + & \cdot \\ \cdot & + \end{array}\right]$ | $\left[\begin{array}{cc}1&.\\.&1\end{array}\right]$ | $\left[\begin{array}{cc} \cdot & \cdot \\ \cdot & + \end{array}\right]$ | $\left[\begin{array}{cc} 1 & . \\ . & 1 \end{array}\right]$ |
| Venezuela | currency banking | $\left[\begin{array}{cc} + & . \\ . & + \end{array}\right]$ | $\left[\begin{array}{cc} 1 & + \\ + & 1 \end{array}\right]$ | + . . + | $\left[\begin{array}{cc} 1 & + \\ + & 1 \end{array}\right]$ | $\left[\begin{array}{cc} \cdot & \cdot \\ \cdot & + \end{array}\right]$ | $\left[\begin{array}{cc} 1 & + \\ + & 1 \end{array}\right]$ |

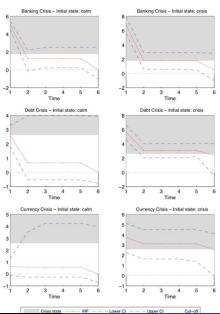
Application- Bivariate vs trivariate DPM 2

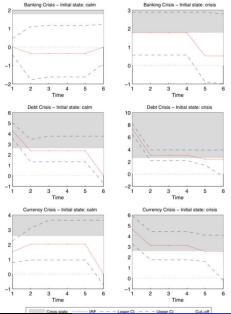
Table 4 – Trivariate Analysis

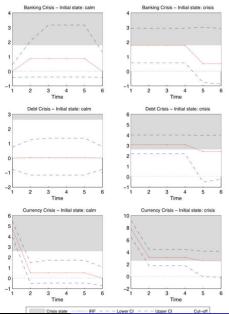
| | | 3 m | onths | 6 months | | |
|--------------|----------------------------------|---|---|---|---|--|
| Country | | θ | Ω | θ | Ω | |
| Ecuador | currency banking sovereign | + . + . + | 1 . . 1 . 1 | | 1 . . 1 . 1 | |
| South Africa | currency banking sovereign | $\left[\begin{array}{ccc} + & \cdot & \cdot \\ \cdot & \cdot & \cdot \\ \cdot & \cdot & + \end{array}\right]$ | $\left[\begin{array}{ccc} 1 & . & + \\ . & 1 & . \\ + & . & 1 \end{array}\right]$ | $\left[\begin{array}{ccc} + & \cdot & \cdot \\ \cdot & \cdot & \cdot \\ \cdot & \cdot & + \end{array}\right]$ | $\left[\begin{array}{ccc} 1 & . & + \\ . & 1 & . \\ + & . & 1 \end{array}\right]$ | |

Note: Two different lags of the dependent variable are used, namely 3 and 6 months. ' θ ' stands for the parameters of the lagger crisis variables, while Ω represents the variance-covariance matrix. $\Lambda'+'/^{1/2}$ sign means that the coefficient is significant and positive/ negative, while a '.' indicates its non-significance. For example, in the case of Ecuador, 3 months, sovereign debt crise have a positive and significative impact on the probability of occurrence of currency crises.









Conclusion

- 1. This paper develops a multivariate dynamic probit model for financial crisis model,
- 2. It allows to model the potential mutation of a crisis into another one.
- 3. We also propose an exact Maximum Likelihood estimation.
- 4. MDP shows better in-sample properties as it takes into account crisis mutation.
- → Powerful tool for real life implementation...